

# TESTS OF HYPOTHESES

## BIOS 662

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## Basic Approach

1. Set up a hypothesis
2. Collect data
3. Infer from the data whether hyp is plausible

- Examples:

Is BP the same in diabetics and non-diabetics?

Will folic acid supplementation reduce the risk of stroke?

## Null Hypothesis: $H_0$

- Null hypothesis  $H_0$ : hypothesis to be tested
- Example null hyp for folic acid study:  
the incidence of stroke will be the same in those taking folic acid supplements and those not taking folic acid supplements
- See Note 4.16 text: typically  $H_0$ 
  - (i) prevailing view or straw man
  - (ii) most parsimonious hypothesis

## Null and Alternative

- In a test of a hyp, we are testing whether some population parameter has a particular value
- For example,

$$H_0 : \theta = \theta_0$$

where  $\theta_0$  is a known constant

- The [alternative hypothesis](#) is complement of null hyp

$$H_A : \theta \neq \theta_0$$

## Test Statistic

- Once the data are collected, we will compute a *test statistic* related to  $\theta$ , say  $S(\hat{\theta})$
- $S(\hat{\theta})$  is a random variable, since it is computed from a sample
- $S(\hat{\theta})$  will have a particular probability distribution under the assumption  $H_0$ , say  $F_0[S(\hat{\theta})]$

## Test Statistic

- Under  $F_0$ , we compute the probability that we would observe  $S(\hat{\theta})$  or a value more extreme than  $S(\hat{\theta})$  if the null  $H_0$  was true
- If this probability is large, the data are consistent with  $H_0$
- If this probability is small, two possibilities:
  1. An unlikely event has occurred
  2.  $H_0$  is not true

## Interpretation

- Usually if the probability is small, we conclude  $H_0$  is not true; i.e., we “reject”  $H_0$
- If the probability is large, we have not proved  $H_0$ . We say that “we failed to reject  $H_0$ ”
- We can never prove  $H_0$  is true!
- Also: don’t “accept the alternative”

## Significance Level

- How do we decide if the probability is too small?
- Prior to seeing the data, we select a value  $\alpha$  such that:  
if the computed probability is less than or equal to  $\alpha$ ,  
we reject  $H_0$
- $\alpha$  is known as *significance level*



## Critical Region and Value

- We have a statistic  $S(\hat{\theta})$  with distribution  $F_0$  under the null hyp
- We specify  $\alpha$  and under  $F_0$  determine a *critical region* or *rejection region*  $C_\alpha$  such that

$$\Pr[S(\hat{\theta}) \in C_\alpha | H_0] = \alpha$$

- Values at the boundaries of  $C_\alpha$  are call *critical values*

## Critical Region and Value

- From the data we compute  $S(\hat{\theta})$
- If  $S(\hat{\theta}) \in C_\alpha$ , we reject  $H_0$
- If  $S(\hat{\theta}) \notin C_\alpha$ , the data are consistent with  $H_0$  and we do not reject  $H_0$

## Tests of Hypotheses: Seven Steps

1. Design study
2. Establish null hyp
3. Determine test statistic to be employed
4. Choose significance level  $\alpha$  and establish  $C_\alpha$
5. Carry out study and collect data
6. Compute statistic from data
7. If statistic is in  $C_\alpha$ , reject  $H_0$

## Example

- Does calcium supplementation **affect** blood pressure in African Americans with high blood pressure?
- Study: take 10 AA men with hypertension; measure BP; ask them to take calcium tablets for 3 weeks and re-measure their BP

## Example

- Let  $\theta$  denote the mean BP change after 3 weeks
- Hypotheses

$$H_0 : \theta = 0 \text{ vs } H_A : \theta \neq 0$$

- Let  $Y_i = \text{BP at 3 weeks} - \text{BP at baseline}$  for the  $i^{\text{th}}$  individual in the study,  $i = 1, \dots, 10$
- $\hat{\theta} = \bar{Y}$

## Example

- Intuition: want to reject  $H_0$  if  $\bar{Y}$  is far from  $\theta_0 = 0$ , i.e., if

$$|\bar{Y}| > c$$

for some constant  $c$

- In particular, want  $c$  such that

$$\Pr[|\bar{Y}| > c | H_0] = \alpha$$

## Example

- Equivalently, want  $c$  such that

$$\Pr \left[ \left| \frac{\bar{Y}}{s/\sqrt{n}} \right| \geq \frac{c}{s/\sqrt{n}} \mid H_0 \right] = \alpha$$

- Assuming  $Y_i$  are iid  $N(\theta, \sigma^2)$ , under  $H_0$

$$\frac{\bar{Y}}{s/\sqrt{n}} \sim t_{n-1}$$

- Thus choose  $c$  such that

$$\frac{c}{s/\sqrt{n}} = t_{n-1, 1-\alpha/2}$$

## Example

- So we reject  $H_0$  if

$$|\bar{Y}| > c = t_{n-1, 1-\alpha/2} s / \sqrt{n}$$

i.e. if

$$\frac{|\bar{Y}|}{s/\sqrt{n}} > t_{n-1, 1-\alpha/2}$$

- Equivalently

$$C_\alpha = \{T : |T| > t_{n-1, 1-\alpha/2}\}$$

where

$$T = \frac{\bar{Y}}{s/\sqrt{n}}$$



## Example

- Returning to calcium example,

$$S(\hat{\theta}) = S(\bar{Y}) = \frac{\bar{Y} - \theta_0}{s/\sqrt{n}} \sim t_9$$

- If  $\alpha = 0.05$ , the critical region is

$$C_{.05} = \{T : |T| > t_{9,.975} = 2.26\}$$

where

$$T = \frac{\bar{Y} - 0}{s/\sqrt{10}}$$

# Example

Calcium supplementation in African-American men

	treatment	before	after	diff
1.	calcium	107	100	-7
2.	calcium	110	114	4
3.	calcium	123	105	-18
4.	calcium	129	112	-17
5.	calcium	112	115	3
6.	calcium	111	116	5
7.	calcium	107	106	-1
8.	calcium	112	102	-10
9.	calcium	136	125	-11
10.	calcium	102	104	2

## Example in R

```
> x <- c(-7,4,-18,-17,3,5,-1,-10,-11,2)
> se <- sd(x)/sqrt(length(x))
> mean(x)/se
[1] -1.808411

> t.test(x)
```

### One Sample t-test

```
data:  x
t = -1.8084, df = 9, p-value = 0.104
```

# Example in SAS Proc Univariate

The UNIVARIATE Procedure

Variable: x

## Moments

N	10	Sum Weights	10
Mean	-5	Sum Observations	-50
Std Deviation	8.74325137	Variance	76.44444444

## Tests for Location: Mu0=0

Test	-Statistic-	-----p Value-----
Student's t	t -1.80841	Pr >  t  0.1040

## Example

- Since the observed  $t = -1.8084$  is not in the critical region, we do not reject  $H_0$
- The data are consistent with no effect of calcium on BP

## Types of Errors

		Nature	
		$H_0$ true	$H_A$ true
Decision	Do not reject $H_0$	✓	Type II
	Reject $H_0$	Type I	✓

- Type I error: Reject  $H_0$  when  $H_0$  true  
i.e. false positive
- Type II error: do not reject  $H_0$  when  $H_A$  true  
i.e. false negative

## Types of Errors

- Type I error

$$\alpha = \Pr[S(\hat{\theta}) \in C_\alpha | H_0]$$

- Type II error

$$\beta = \Pr[S(\hat{\theta}) \notin C_\alpha | H_A]$$

- Power

$$1 - \beta = \Pr[S(\hat{\theta}) \in C_\alpha | H_A]$$

i.e., prob reject  $H_0$  when  $H_A$  is true

## Power

- Recall  $H_A : \theta \neq \theta_0$
- Power:  $\Pr[S(\hat{\theta}) \in C_\alpha | H_A]$
- Power depends on the value of  $\theta$

$$\Pr[S(\hat{\theta}) \in C_\alpha | \theta] \equiv P(\theta)$$

- Note

$$P(\theta_0) = \alpha$$



## Alternative Hypotheses

- Different possible alternatives

$$H_A : \theta \neq \theta_0 \text{ (two-sided)}$$

$$H_A : \theta > \theta_0 \text{ (one-sided)}$$

$$H_A : \theta < \theta_0 \text{ (one-sided)}$$

## Alternative Hypotheses

- In the AA example,

$$H_A : \theta \neq 0; C_{.05} = \{T : |T| > 2.26\}$$

- If we took a 1-sided alternative,

$$H_A : \theta < 0; C_{.05} = \{T : T < -1.83\}$$

- Since  $T = -1.8084$ , we do not reject  $H_0$  in either setting

## Alternative Hypotheses

- 2-sided test answers: Does calcium **change** BP?
- 1-sided test answers: Does calcium **lower** BP?
- If we apply 1-sided test and get  $T = 2.5$ , we do not reject  $H_0$  because our test did not ask if calcium raised or changed BP
- We must choose the alternative hyp **before seeing the data**
- Friedman et al (p. 98) “In general, two-sided tests should be used unless there is strong justification for expecting a difference in only one direction”

## P-value

- Def 4.24 The *p-value* is the smallest significance level  $\alpha$  for which the observed data indicate the null hypothesis should be rejected
- Probability of obtaining test statistic as unlikely or more unlikely than the observed test statistic if the null hypothesis is true

## Calcium Example Revisited

- Recall  $T \sim t_9$
- p-value for 2-sided test =  $2 \Pr[T < -1.804] = 0.1040$
- p-value for 1-sided test =  $\Pr[T < -1.804] = 0.0520$
- R

```
> 2*pt(-1.8084,9)
```

```
[1] 0.1039981
```

```
> pt(-1.8084,9)
```

```
[1] 0.05199907
```