

1. We identify the kernel function  $h(X_1, X_2) = |X_1 - X_2|$ . Apply the CLT for the U-statistics and after the calculation under the uniform distribution, we obtain

$$\sqrt{n} \left\{ \binom{n}{2}^{-1} \sum_{i < j} |X_i - X_j| - \frac{1}{3} \right\} \rightarrow N\left(0, \frac{1}{90}\right).$$

2. Clearly,  $E[|X_n|] < \infty$ . Let  $\mathcal{F}_n = \sigma(Y_1, \dots, Y_n)$ . Then  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$ . Since

$$X_{n+1} = \left( \sum_{k=1}^n Y_k \right)^2 + 2 \left( \sum_{k=1}^n Y_k \right) Y_{n+1} + Y_{n+1}^2 - (n+1)\sigma^2,$$

we have

$$\begin{aligned} E[X_{n+1} | \mathcal{F}_n] &= \left( \sum_{k=1}^n Y_k \right)^2 + 2 \left( \sum_{k=1}^n Y_k \right) E[Y_{n+1} | Y_1, \dots, Y_n] + E[Y_{n+1}^2 | Y_1, \dots, Y_n] - (n+1)\sigma^2 \\ &= \left( \sum_{k=1}^n Y_k \right)^2 + 2 \left( \sum_{k=1}^n Y_k \right) E[Y_{n+1}] + E[Y_{n+1}^2] - (n+1)\sigma^2 = \left( \sum_{k=1}^n Y_k \right)^2 - n\sigma^2 = X_n. \end{aligned}$$

Therefore,  $\{X_n\}$  is a martingale.

3. (a) Clearly,  $E[T_n] = \sum_{i=1}^n \omega_{ni} \mu = \mu$  and

$$Var(T_n) = \sum_{i=1}^n \omega_{ni}^2 \sigma_i^2.$$

We aim to minimize  $Var(T_n)$  under constraint  $\sum_{i=1}^n \omega_{ni} = 1$  and  $\omega_{ni} \geq 0$ . This can be done using the Lagrange multiplier by solving

$$\sigma_i^2 \omega_{ni} - \lambda = 0.$$

This yields  $\omega_{ni} = 1/(\lambda \sigma_i^2)$ . Since  $\sum_{i=1}^n \omega_{ni} = 1$ , we obtain the optimal weight.

- (b) When  $\omega = \omega^{opt}$ ,

$$Var(T) = \left\{ \sum_{i=1}^n 1/\sigma_i^2 \right\}^{-1}.$$

Since  $Var(T) \rightarrow 0$ , by the Chebyshev's inequality,

$$P(|T_n - \mu| > \epsilon) \leq \epsilon^{-2} Var(T_n) \rightarrow 0.$$

Thus,  $T_n \rightarrow_p \mu$ .

(c) When  $X_i = \mu + \sigma_i \epsilon_i$ ,

$$\sqrt{\sum_{i=1}^n (1/\sigma_i^2)} (T_n - \mu) = \sum_{i=1}^n \frac{1/\sigma_i}{\sqrt{\sum_{i=1}^n (1/\sigma_i^2)}} \epsilon_i.$$

We verify the Lindeberg conditions. Clearly, the above variance is 1. Furthermore, for any  $\delta > 0$ ,

$$\begin{aligned} & \sum_{i=1}^n E \left[ \left( \frac{1/\sigma_i}{\sqrt{\sum_{i=1}^n (1/\sigma_i^2)}} \epsilon_i \right)^2 I \left( \frac{1/\sigma_i}{\sqrt{\sum_{i=1}^n (1/\sigma_i^2)}} |\epsilon_i| > \delta \right) \right] \\ & \leq \sum_{i=1}^n \frac{1/\sigma_i^2}{\sum_{i=1}^n (1/\sigma_i^2)} E \left[ \epsilon_1^2 I \left( |\epsilon_1|^2 > \delta^2 / [\max(1/\sigma_i^2) / \sum_{j=1}^n (1/\sigma_j^2)] \right) \right] \\ & \leq E \left[ \epsilon_1^2 I \left( |\epsilon_1|^2 > \delta^2 / [\max(1/\sigma_i^2) / \sum_{j=1}^n (1/\sigma_j^2)] \right) \right] \rightarrow 0. \end{aligned}$$

Hence, the result follows from the Lindeberg-Feller CLT.

(d)

$$\text{Var}(T_n) / \text{Var}(\bar{X}_n) = \frac{n^2}{\sum_{j=1}^n (1/\sigma_j^2) \sum_{j=1}^n \sigma_j^2}.$$

The following table gives the ratios for difference choices of  $r$  and  $n$ :

$r$	$n = 5$	$n = 10$	$n = 20$	$n = 50$	$n = 100$	$n = \infty$
0.25	5.503909e-02	2.145771e-04	8.185452e-10	4.437343e-27	1.400178e-56	0
0.5	4.162331e-01	4.892363e-02	1.907352e-04	1.110223e-12	3.944305e-27	0
0.75	8.498904e-01	5.269579e-01	1.063807e-01	1.179838e-04	2.672668e-10	0