## Solution to BIOS 760 Midterm Exam I (Fall, 2016)

1. (a)  $(\alpha' X, \beta' X)$  follows a bivariate normal distribution with mean zeros and covariance

$$\begin{pmatrix} \alpha' \Sigma \alpha & \alpha' \Sigma \beta \\ \beta' \Sigma \alpha & \beta' \Sigma \beta \end{pmatrix}.$$

(b) Let  $Z = \alpha' X - (\alpha' \Sigma \beta) / (\beta' \Sigma \beta) \beta' X$ . Then Z is independent of  $\beta' X$  since  $cov(\beta' X, Z) = 0$ . Clearly,

$$\alpha' X = a\beta' X + Z, \quad a = (\alpha' \Sigma \beta)/(\beta' \Sigma \beta).$$

(c)  $\alpha' X/|\beta' X| = a + Z/|\beta' X|$ . Since  $Z \sim \sqrt{\alpha' \Sigma \alpha - (\alpha' \Sigma \beta)^2/(\beta' \Sigma \beta)} N(0, 1)$  and  $|\beta' X| \sim \sqrt{\beta' \Sigma \beta} \chi^2(1)$ , we obtain

$$\alpha' X / \beta' X \sim a + \frac{\sqrt{\alpha' \Sigma \alpha - (\alpha' \Sigma \beta)^2 / (\beta' \Sigma \beta)}}{\sqrt{\beta' \Sigma \beta}} t(1).$$

- 2. (a) This is because  $(\mu \times \lambda_F)(\Omega_1 \times \Omega_2) = \mu(\Omega_1) \times \lambda_F(\Omega_2) = 1.$ 
  - (b)  $\{X \le x\} = \bigcup_{k=1}^{\infty} \{(w_1, w_2) : w_1 = k, I(w_2 \in [k, k+1)) \le x\} = \bigcup_k \{w_1 : w_1 = k\} \times \{w_2 : I(w_2 \in [k, k+1)) \le x\} \in 2^{\Omega_1} \times \mathcal{B}.$  Thus, X is measurable.
  - (c) By the Fubini theorem,

$$\begin{split} & \int_{\Omega_1 \times \Omega_2} X(\omega_1, \omega_2) d(\mu \times \lambda_F) \\ &= \int_{\Omega_1} \int_{\Omega_2} I(w_2 :\in [w_1, w_1 + 1)) d\lambda_F(w_2) d\mu(w_1) \\ &= \int_{\Omega_1} \int_{w_1}^{w_1 + 1} dF(w_2) d\mu(w_1) \\ &= \int_{\Omega_1} \left\{ e^{-w_1} - e^{-w_1 - 1} \right\} d\mu(w_1) \\ &= \sum_{k=1}^{\infty} 2^{-k} \left\{ e^{-k} - e^{-k - 1} \right\} \\ &= \frac{e^{-1}/2}{1 - e^{-1/2}} (1 - e^{-1}). \end{split}$$

(d) This is clear from  $P(Y \le y, Z \le z) = \mu(Y \le y)\lambda_F(Z \le z)$ .

(e) The measure induced by Y is dominated by the counting measure on  $\{1, 2, ..\}$  with density  $P(Y = k) = 2^{-k}$ . The measured induced by Z is dominated by  $\lambda$  with density  $e^{-x}I(x > 0)$ .

(f) 
$$E[X] = E[E[I(Z \in [Y, Y+1))|Y]] = E[e^{-Y} - e^{-(Y+1)}] = \frac{e^{-1/2}}{1 - e^{-1/2}}(1 - e^{-1}).$$