Bayesian QTL Mapping

Fei Zou Department of Biostatistics Gillings School of Global Public Health University of North Carolina-Chapel Hill fzou@bios.unc.edu

- One gene one trait
 - very unlikely
- Most traits have a significant environmental exposure component
- The vast majority of biological traits are caused by complex polygenic interactions

▲ロト ▲帰ト ▲ヨト ▲ヨト 三日 - の々ぐ

• also context dependent

Single QTL Mapping

- Single marker analysis (Sax, 1923 Genetics)
- Interval mapping: Lander & Botstein (1989, Genetics)

Multiple QTL mapping

• Composite interval mapping (Zeng 1993 PNAS, 1994 Genetics; Jansen & Stam, 1994 Genetics)

- Multiple interval mapping (Kao et al., 1999 Genetics)
- Bayesian analysis (Satagopan et al., 1997 Genetics)

 Most complicated traits are caused by multiple (potentially interacting) genes, which also interact with environment stimuli

- Single QTL interval mapping
 - Ghost QTL (Lander & Botstein 1989)
 - Low power

- Bayesian methods (Stephens and Fisch 1998 Biometrics; Sillanpaa and Arjas 1998 Genetics; Yi and Xu 2002 Genetic Research, and Yi et al. 2003 Genetics): treat the number of QTLs as a parameter by using reversible jump Markov chain Monte Carlo (MCMC) of Green (1995 Biometrika)
 - change of dimensionality, the acceptance probability for such dimension change, which in practice, may not be handled correctly (Ven 2004 Genetics)

- Alternative, multiple QTL mapping can be viewed as a variable selection problem
 - Forward and step-wise selection procedures (Broman and Speed 2002 JRSSB)
 - LASSO, etc
 - Bayesian QTL mapping
 - Xu (2003 Genetics), Wang et al (2005 Genetics) Huang et al (2007 Genetics): Bayesian shrinkage
 - Yi et al (2003 Genetics): stochastic search variable selection (SSVS) of George and McCulloch (1993 JASA)
 - Yi (2004 Genetics): composite model space of Godsill (2001 J. Comp. Graph. Stat)
 - Software: R/qtlbim by Yi's group

- Limitations of existing QTL mapping methods
 - do not model covariates at all or only model covariate effect linearly

• do not model interactions at all or model only lower order interactions, such as two way interactions

• The multiple QTL mapping is a very large variable selection problem: for *p* potential genes, with *p* being in the hundreds or thousands, there are 2^{*p*} possible main effect models, $\begin{pmatrix} p \\ 2 \end{pmatrix} possible two-way interactions and 2 \begin{pmatrix} p \\ k \end{pmatrix} possible higher order (k > 2) interactions.$

<ロト 4 回 ト 4 回 ト 4 回 ト 回 の Q (O)</p>

Semiparmetric QTL Mapping: Non-linear Covariate Effect

- Goal: allow arbitrary covariate effect in QTL Mapping model
- Semiparametric model:

$$y_i = \sum_j \beta_j x_{ij} + \eta(t_i) + e_i, \quad i = 1, \cdots n \quad \text{with } e_i \sim N(0, \sigma^2)$$

 Function η is unspecified and t_i is a scalar or a vector of covariates

• Deifne
$$\tilde{x}_j = (x_{1j}, \cdots, x_{nj})'$$
 and $\mathbf{x}_i = (x_{i1}, \cdots, x_{ip})'$

Existing Semi/non-parametric Methods

- Dirichlet process (Muller et al. 1996)
- Splines (Smith and Kohn 1996; Denison et al. 1998 and DiMatteo et al. 2001)
- Wavelets (Abramovich et al. 1998 JRSSB)
- Kernel models (Feng et al 2007)
- Gaussian process (Neal 1997; 1996)
 - Gaussian process priors have a large support in the space of all smooth functions through an appropriate choice of covariance kernel.
 - Gaussian process is flexible for curve estimation because of their flexible sample path shapes
 - Gaussian process related to smoothing spline somehow (Wahba 1978 JRSSB)

Semiparmetric QTL Mapping: Non-linear Covariate Effect

Model

 $y_i = \sum_j \beta_j x_{ij} + \eta(t_i) + e_i, \quad i = 1, \cdots n \quad \text{with } e_i \sim N(0, \sigma^2)$

• The unobserved variables are: function η , error variane σ^2 and the QTL effects $\beta = \{\beta_j\}$

• A Gaussian process where all possible finite dimensional distributions $\eta = (\eta(t_1), \ldots, \eta(t_n))'$ follow a multivariate normal with $E(\eta(t_i)) = \mu(t_i)$ and $cov(\eta(t_i), \eta(t_j)) = \sigma(t_i, t_j)$ where

$$\mu(t; lpha) = lpha_1 f_1(t) + \dots + lpha_l f_l(t)$$
 and
 $\sigma(t_i, t_j) = \frac{1}{ au} \exp(-
ho(t_i - t_j)^2)$

• τ controls the smoothness of η : when $\tau \to 0$, the posterior mean of η almost interpolates the data while centered around the prior mean function if $\tau \to \infty$.

- On $\beta = \{\beta_j\}$: follow SSVS idea such that $P(\gamma_j = 1) = 1 - P(\gamma_j = 0) = p_j$ $\beta_i \mid \gamma_j \sim (1 - \gamma_j)N(0, \sigma_0^2) + r_jN(0, \gamma_0^2)$
- we also put hyper priors on τ and $\alpha = (\alpha_1, \dots, \alpha_l)$ as $\tau \sim \text{inverse-Gamma}(a_{\tau}, b_{\tau})$ and $\alpha \sim N(\alpha_0, \Gamma)$

Conditional Probability

- Define $\Sigma = \tau Var(\eta)$, i.e., prior of η is $N(\mu, \Sigma/\tau)$.
- Define **M** as a $k \times l$ matrix with the (i, j)th element equal to $\mu_i(t_i)$.
- Then the conditional posterior distributions are

•
$$\beta_j \mid \mathbf{y}, \boldsymbol{\theta}_{-\beta_j} \sim N(\hat{\beta}_j, \frac{\sigma^2}{\mathbf{x}'_j \mathbf{x}_j + \sigma^2 / \sigma_j^2})$$
 [exercise: find $\hat{\beta}_j$]
• $p(\gamma_j = 1 \mid \boldsymbol{\theta}_{-\gamma_j}) = \frac{c_j}{c_j + \frac{p_j}{1 - p_j} \exp\left\{\frac{\beta_j^2}{2\sigma_1^2}\left(1 - \frac{1}{c_j^2}\right)\right\}}$
• $\eta \mid \mathbf{y}, \alpha, \beta, \tau \sim N(\mu^*, \boldsymbol{\Sigma}^*)$
•

where θ is the vector of all unknown parameters, $\sigma_i = c_i^{\gamma_j} \sigma_0$, $\mu^{\star} = \mathbf{\Sigma}^{\star} \mathbf{D} (\mathbf{U} - \mu) + \mu, \ \mathbf{\Sigma}^{\star} = (\mathbf{D} + \mathbf{\Sigma}^{-1})^{-1},$ $\mathbf{D} = diag(d_1, \dots, d_k)/\sigma^2$ with $d_i = \#$ of subjects that have covariates equal to t_i and $\mathbf{U} = \{u_i\}$ with u_i = average of $y_i - \sum_i \beta_j x_{ij}$ for those samples whose covariate equals t_i .

Bayesian Model: Simulation Results

• See lecture4 ppt file



- Goal: map multiple potentially interacting QTLs without specifically model all potential main and higher order interaction effects
- Semiparametric model:

 $y_i = \eta(x_{i1}, \cdots, x_{ip}) + e_i, \quad i = 1, \cdots, n \text{ with } e_i \sim N(0, \sigma^2)$

- Again function η is unspecified
 - very flexible
 - $\eta(x_{i1}, \cdots, x_{ip}) = x_{i1} + x_{i3}$, or $x_{i1}x_{i3}$ or $x_{i1} + x_{i4}x_{i5}x_{i6}$

 \bullet Again Gaussian process prior is placed on η function such that

- $E(\eta_i) = 0$
- $cov(\eta_i, \eta_k) = \frac{1}{\tau} exp[-\sum_j \rho_j(x_{ij} x_{kj})^2]$ where $\eta_i = \eta(x_{i1}, \cdots, \eta_{ip})$ and let $\eta = \{\eta_i\}$.
- Hyperparameters ρ_j related to length scales $\frac{1}{\sqrt{\rho_j}}$ which characterize the distance in that particular direction over which η is expected to vary significantly.
- When ρ_j = 0, η is expected to be an essentially constant function of that input variable j, which is therefore deemed irrelevant (Mackay 1998).

- The original papers on the Gaussian process (Mackay 1998; Neal 1997) did not view this method as an approach for variable selection and imposed a Gamma prior on the ρ_j parameters. However, ρ_j does provide information about the relevance of any QTL with value near zero indicating an irrelevant QTL.
- For variable selection purpose, we can impose the following mixture priors on ρ_i based on latent variable γ_i:

•
$$P(\gamma_j = 1) = p_j$$

• $\tau_i (= 1/\rho_i) \sim (1 - \gamma_i) Ga(\frac{\alpha_0}{2}, \frac{\alpha_0}{2\alpha_i}) + \gamma_i Ga(\frac{\alpha_1}{2}, \frac{\alpha_1}{2\alpha_i})$

- No closed posterior form for τ_j s and we resort to Metropolis-Hastings algorithm
 - Direct use of MH is not very efficient for our model and it would explore region of hight probability by an very inefficient random walk
 - hybrid MC method was proposed (Neal 1993,1996: Rasmussen 1996; Barber et al 1997) and we adopt this approach
 - hybird approach merges the MH algorithm with sampling techniques called dynamic simulation based on a "energy" function
- Not computationally feasible for GWAS data where millions of genotypes available on thousands of samples
 - deterministic algorithms to replace MCMC sampling, such as conjugate gradient optimization technique for maximum-a-posterior estimates?

Semiparmetric QTL Mapping: Extensions

- A) non-genetic factors, z_i = (z_{i1}, · · · , z_{iq}) can be also inluded into η
 y_i = η(x_{i1}, · · · , x_{ip}, z_{i1}, · · · , z_{iq}) + e_i, i = 1, · · · , n with e_i ~ N(0, σ²)
- B) longitudinal data $y_{ij} = \eta(x_{i1}, \dots, x_{ip}, t_{ij}) + e_{ij}$ with $\mathbf{e} = (e_{i1}, \dots, e_{i,k_i}) \sim N(0, \mathbf{\Sigma}_i)$. We have considered cases where
 - Σ_i is known up to certain parameters
 - Σ₁ is unknown and modelled vai the deomposition method of Chen and Dunson (2003, Biometrics)

(日) (同) (三) (三) (三) (○) (○)